Lecture 8 · 7.5.2021 · 10:15-12:00 via Zoom

## **Conditional expectation (end)**

**Example.** Let  $(X_{n,m})_{n,m\geqslant 0}$  a double sequence of i.i.d. r.v. with values in  $\mathbb{N}_{\geqslant 0} = \{0,1,2,\ldots\}$  (in particular  $\mathbb{P}(X_{1,1} < \infty) = 1$ ). Let  $Z_0 = 1$  and then I define recursively

$$Z_n = X_{n,1} + \cdots + X_{n,Z_{n-1}},$$

for  $n \ge 1$ . Note that  $Z_n$  is a integer valued random variable.

The r.v.s  $(Z_n)_{n\geqslant 0}$  model the evolution of the size of population of individuals where at every step n the m-th individual give rise to  $X_{n,m} \sim X_{1,1}$  individuals which go the next generation.

We want to compute the generating function  $f_n(\theta) := \mathbb{E}[\theta^{Z_n}]$  of  $Z_n$  for every  $\theta \in (0,1)$ .

Note that we can more rigorously define

$$Z_n = \sum_{m>1} \mathbb{1}_{Z_{n-1} \geqslant m} X_{n,m},$$

and observe that  $Z_0$  it is a.s. finite and if  $Z_{n-1}$  is a.s. finite then the above sum is also made a.s. of finitely many summands and therefore  $Z_n$  is a.s. finite if  $\mathbb{P}(X_{1,1} < \infty) = 1$ , i.e. this proves that  $\mathbb{P}(Z_n < \infty) = 1$  for all  $n \ge 0$  by induction, assuming  $\mathbb{P}(X_{1,1} < \infty) = 1$ .

We know that  $f_0(\theta) = \mathbb{E}[\theta^{Z_0}] = \theta$ . Let us call  $f(\theta) := \mathbb{E}[\theta^{X_{1,1}}]$ . How we compute

$$f_n(\theta) = \mathbb{E}[\theta^{Z_n}] = \mathbb{E}[\theta^{X_{n,1}+\cdots+X_{n,Z_{n-1}}}]$$
?

We note that  $(X_{n,m})_{m\geqslant 1}$  is independent of  $Z_{n-1} \in \sigma(X_{k,m}: 0 \le k \le n-1, m\geqslant 1)$  since the family  $(X_{n,m})_{n,m\geqslant 0}$  is iid. In this case we can condition the averange on the value of  $Z_{n-1}$ : by the theorem we proved in the last lecture (Prop 14 in Note 2)

$$\mathbb{E}[\theta^{X_{n,1}+\cdots+X_{n,Z_{n-1}}}] = \mathbb{E}[\mathbb{E}[\theta^{X_{n,1}+\cdots+X_{n,Z_{n-1}}}|Z_{n-1}]] = \mathbb{E}[\varphi(Z_{n-1})]$$

with

$$\varphi(z) = \mathbb{E}[\theta^{X_{n,1} + \dots + X_{n,z}}], \qquad z = 0, 1, 2, \dots$$

But now we can compute this easily since cond. exp. has disappeared thanks to independence, for  $z \ge 0$ 

$$\varphi(z) = \mathbb{E}[\theta^{X_{n,1}}] \cdots \mathbb{E}[\theta^{X_{n,z}}] = (\mathbb{E}[\theta^{X_{1,1}}])^z = (f(\theta))^z$$

which means that

$$f_n(\theta) = \mathbb{E}[\theta^{X_{n,1}+\cdots+X_{n,Z_{n-1}}}] = \mathbb{E}[\varphi(Z_{n-1})] = \mathbb{E}[(f(\theta))^{Z_{n-1}}] = f_{n-1}(f(\theta)).$$

Therefore we have shown that  $f_n$  solve the recursive equation

$$f_0(\theta) = \theta, \qquad f_n(\theta) = f_{n-1}(f(\theta))$$

which has unique solution  $f_n(\theta) = f^{\circ n}(\theta)$ . This is very useful if one would like to understand what happens to  $Z_n$  as  $n \to \infty$ , i.e. how the population behave on long time. Intersting question: does it become extinct with probability 1. Note that if  $Z_n = 0$  then  $Z_k = 0$  for all  $k \ge n$ .

## Regular conditional probabilities

Given a  $\sigma$ -algebra  $\mathscr{G} \subseteq \mathscr{F}$  we can consider the family of r.v.

$$A \in \mathcal{F} \mapsto \mathbb{P}(A|\mathcal{G}) \coloneqq \mathbb{E}[\mathbb{1}_A|\mathcal{G}] \in L^{\infty}(\Omega,\mathcal{G},\mathbb{P}) \subseteq L^{\infty}(\Omega,\mathcal{F},\mathbb{P})$$

These r.v. satisfy **almost surely** the following equalities:

a) 
$$\mathbb{P}(\emptyset|\mathscr{G}) = 0$$
,  $\mathbb{P}(A^c|\mathscr{G}) = 1 - \mathbb{P}(A|\mathscr{G})$ 

b) 
$$\mathbb{P}(\bigcup_n A_n | \mathcal{G}) = \sum_n \mathbb{P}(A_n | \mathcal{G})$$
 for a family  $(A_n)_{n \ge 1} \subseteq \mathcal{F}$  of pairwise disjoint events.

These relations shows that the map  $A \in \mathcal{F} \mapsto \mathbb{P}(A|\mathcal{G})$  behaves like a probability measure. So we would like to think to  $\mathbb{P}(\cdot|\mathcal{G})$  as a random probability measure  $\mathbb{P}_{\mathcal{G}}$ , i.e. something like

$$\mathbb{P}_{\mathscr{C}}: \omega \in \Omega \mapsto \mathbb{P}_{\mathscr{C}}(\omega) \in \Pi(\Omega, \mathscr{F})$$

where  $\Pi(\Omega, \mathcal{F})$  is the set of all probability measures on  $(\Omega, \mathcal{F})$ . This is in this generality not possible, because the properties a),b) are true only a.s., that is for any choice of A or of  $(A_n)_n$  one has different exceptional set  $\mathcal{N}_A$ ,  $\mathcal{N}_{(A_n)_n}$  in which the property is not satisfied and unfortunately one cannot construct an exceptional universal measurable set valid for all the possible choices of A,  $(A_n)$ , because we have uncountably many choices here.

In some situations however this is possible. In this case we say that the family  $(\mathbb{P}(A|\mathcal{G}))_{A\in\mathcal{F}}$  admits a **regular conditional version**, or that we have a **regular conditional probability** for  $\mathbb{P}$  given  $\mathcal{G}$ . This means that there exists a map

$$\mathbb{P}_{\mathscr{C}}: \Omega \to \Pi(\Omega, \mathscr{F}) \subseteq (\Omega \times \mathscr{F} \to [0, 1])$$

such that for all  $A \in \mathcal{F}$  it holds

$$\mathbb{P}_{\mathscr{G}}(\omega, A) = \mathbb{P}(A|\mathscr{G})(\omega), \qquad \mathbb{P} - a.e. \omega \in \Omega.$$

In case we have a regular conditional probability, then for any  $X \in L^1(\Omega, \mathcal{F}, \mathbb{P})$  we can express the conditional expectation wrt.  $\mathcal{G}$  as an integral:

$$\mathbb{E}[X|\mathcal{G}](\omega) = \int X(\omega') \mathbb{P}_{\mathcal{G}}(\omega, d\omega'), \qquad \mathbb{P} - a.e. \, \omega \in \Omega.$$

(this is not difficult to prove from the definition of reg. cond. prob. and cond. exp.)

Regular conditional probabilities are guaranteed to exist when  $(\Omega, \mathcal{F})$  is a **Polish space**: i.e. complete, metrisable topological space endowed with the Borel  $\sigma$ -algebra. Example:  $\mathbb{R}$ ,  $\mathbb{R}^n$ ,  $\mathbb{R}^\mathbb{N}$ ,  $C([0,1], \mathbb{R}^n)$ .

## **Filtrations & stopping times**

**Definition.** A (discrete time) stochastic process is just a family  $(X_n)_{n\geq 0}$  of random variables indexed by  $\mathbb{N}_{\geq 0}$  (or  $\mathbb{N}_{\geq 1}$ ).

We think to the index n at time and to the sequence  $X_1, X_2, ...$  as the description of random phenoma which evolves in time. Time take with it a notion of "past", "present" and "future".

This is encoded in the notion of *filtration*:

**Definition.** A filtration  $(\mathcal{F}_n)_{n\geqslant 0}$  is an increasing family of sub- $\sigma$ -algebras of  $\mathcal{F}$ , i.e.

$$\mathcal{F}_n \subseteq \mathcal{F}_{n+1}$$

for all  $n \ge 0$ . We let  $\mathcal{F}_{\infty} = \sigma(\mathcal{F}_n; n \ge 0)$ , that is the smallest  $\sigma$ -algebra which contains all the  $(\mathcal{F}_n)_n$ .

A filtration represents the flow of time, in the sense that  $\mathcal{F}_n$  is the information I dispose at time n.

**Example.** If  $X = (X_n)_{n \ge 0}$  is a stochastic process then we can always consider its *natural* filtration  $(\mathcal{F}_n^X)_{n \ge 0}$ 

$$\mathcal{F}_n^X = \sigma(X_0, X_1, \dots, X_n).$$

It is easy to check that indeed  $\mathscr{F}_n^X \subseteq \mathscr{F}_{n+1}^X$ . This filtration encode the information given by the observation of the process X as time passes.

**Example.** Let  $\Omega = (0, 1]$  and define

$$\mathcal{F}_n = \{ (k/2^n, (k+1)/2^n) : k = 0, \dots, 2^n - 1 \} \subseteq \mathcal{F} = \mathcal{B}([0, 1])$$

then  $(\mathcal{F}_n)_{n\geqslant 0}$  is a filtration and  $\mathcal{F}_\infty = \mathcal{F}$ . Here *n* represent the precision of our observation of a point in [0,1].

**Definition.** We say that a process  $X = (X_n)_{n \ge 0}$  is adapted to the filtration  $(\mathcal{F}_n)_{n \ge 0}$  iff

$$X_n \in \mathscr{F}_n$$

for all  $n \ge 0$ . A process X is previsible to the filtration  $(\mathcal{F}_n)_{n \ge 0}$  iff

$$X_{n+1} \in \mathscr{F}_n$$

*for all*  $n \ge 0$ .

The natural filtration  $\mathcal{F}^X$  of a process X is the smallest filtration for which the process is adapted.

I want to define now stopping times. A stopping time is a rule to determine how to stop given what happened in the past.

**Definition.** A stopping time  $T: \Omega \to \mathbb{N}^* := \mathbb{N} \cup \{+\infty\}$  for the filtration  $(\mathscr{F}_n)_{n\geqslant 0}$  is a r.v. with values in  $\mathbb{N}^*$  such that

$$\{T \leqslant n\} \in \mathscr{F}_n$$

for all  $n \ge 0$ .

This is equivalent to require that  $\{T = n\} \in \mathcal{F}_n$  for all  $n \ge 0$ .

**Example.** The first time *T* we observe "head" in a repeated launch of a coin, is a stopping time wrt. the natural filtration of this problem. However, the last time *S* I observe "head" is not a stopping time.

The notion of stopping time encode a "fair" stopping rule, i.e. a rule which does not use information from future to make a decision (to stop or no).

**Definition.** The  $\sigma$ -algebra  $\mathcal{F}_T$  of the stopping time T wrt. the filtration  $(\mathcal{F}_n)_{n\geqslant 0}$  is defined as

$$\mathcal{F}_T := \{A \in \mathcal{F} : A \cap \{T \leq n\} \in \mathcal{F}_n \ for \ all \ n \in \mathbb{N}^*\}$$

**Exercise.** Show that  $\mathcal{F}_T$  is a  $\sigma$ -algebra.